

1 Title:

2 **Influence and predictive capacity of climate anomalies**
3 **on daily to decadal extremes in canopy photosynthesis**

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15 Abstract:

16 Significant advances have been made over the past decades in capabilities to simulate
17 diurnal and seasonal variation of leaf-level and canopy-scale photosynthesis in temperate
18 and boreal forests. However, long-term prediction of future forest productivity in a
19 changing climate may be more dependent on how climate and biological anomalies
20 influence extremes in interannual to decadal variability of canopy ecosystem carbon
21 exchanges. These exchanges can differ markedly from leaf level responses, especially owing
22 to the prevalence of long lags in nutrient and water cycling. Until recently, multiple long-
23 term (10+ year) high temporal frequency (daily) observations of canopy exchange were
24 not available to reliably assess this claim. An analysis of one of the longest running North
25 American eddy covariance flux towers reveals that single climate variables do not
26 adequately explain carbon exchange anomalies beyond the seasonal timescale. Daily to
27 weekly lagged anomalies of photosynthesis positively autocorrelate with daily
28 photosynthesis. This effect suggests a negative feedback in photosynthetic response to
29 climate extremes, such as anomalies in evapotranspiration and maximum temperature.
30 Moisture stress in the prior season did inhibit photosynthesis, but mechanisms are difficult
31 to assess. A complex interplay of integrated and lagged productivity and moisture-limiting
32 factors indicate a critical role of seasonal thresholds that limit growing season length and
33 peak productivity. These results lead toward a new conceptual framework for improving
34 earth system models with long-term flux tower observations.

35

36 Keywords: Eddy covariance; canopy photosynthesis; spectral analysis; carbon cycle

37

38 **Introduction**

39 Every year in modern times, photosynthetic organisms on land and in the ocean
40 assimilate around 120,000,000,000,000 kg of carbon dioxide from the atmosphere, a
41 process which drives the entire cycle of biosphere metabolism, production and
42 decomposition (Beer *et al.*, 2010). Variation of photosynthetic rates across space is strongly
43 a function of adaptation of species to climatic, geological, and biological limiting factors of
44 temperature, light, soil nutrients, moisture, disturbance, and competition. These
45 adaptations are often manifested in differences in plant functional form, such as leaf shape,
46 leaf longevity, tree heights, root depths, etc... Similarly, variation of photosynthesis in time
47 is governed by how species in an ecosystem respond and adapt to diurnal, seasonal, and
48 interannual changes in limiting factors.

49 Today, society faces a grand challenge as feedbacks between carbon dioxide uptake
50 by photosynthetic organisms and the climate system are a leading source of uncertainty in
51 the magnitude and severity of future climatic change, on the same order as scenarios of
52 future anthropogenic emissions and aerosol or cloud feedbacks (Booth *et al.*, 2012).
53 Coupled carbon-climate models show a large range of future climate states depending on
54 assumptions built into models about biospheric uptake, particular in the terrestrial
55 biosphere (Friedlingstein *et al.*, 2006). Extreme interannual anomalies in biospheric uptake
56 have been linked to large-scale climate features like El Niño Southern Oscillation (ENSO),
57 and recent increasing trends in the fraction of fossil fuel emissions that remain in the
58 atmosphere point to troubling concerns about the state of the biospheric carbon sink (Le
59 Quéré *et al.*, 2009).

60 Quantifying these variations and improve predictive ecosystem models at the scale
61 of regions to the globe over time periods of days to decades requires careful lab
62 experimentation and long-term field observations (Moorcroft, 2006). Early experiments in
63 the late 1970s and into the 1980s that included careful monitoring of leaf photosynthesis
64 and isotopic discrimination in controlled environments along with theoretical
65 thermodynamic and biochemical arguments led to the first successful representation of leaf
66 photosynthesis through the simplified equations for C3 (and later C4) assimilation as
67 reviewed in Farquhar and Sharkey (1982) and co-occurring development of leaf-
68 atmosphere canopy conductance coupling as reviewed in Collatz *et al* (1991). While major
69 advances have been made on understanding the biochemistry of photosynthesis at the
70 genomic, cellular, and leaf level, many reported in this journal, most of these have not
71 significantly altered these equations and similar formulations that are prevalent in most
72 sophisticated ecosystem models (Schaefer *et al.*, 2012).

73 The reason for the lack of more sophisticated leaf-level photosynthesis models is
74 partly a question of computational resources in that models can not simulate every leaf in
75 an ecosystem, let alone every cell. But a larger source of uncertainty rests in how one goes
76 from the leaf-level model to an ecosystem patch or grid box. Early attempts focused on the
77 issue of scaling of canopy radiative transfer, given that the variation of light through a
78 canopy is the dominant mode of variability of limiting factors within an ecosystem patch to
79 be simulated. Original models include the “big-leaf” representation of the average semi-
80 transparent leaf (e.g., Sellers, 1985), partly owing its success to the ability to characterize
81 vegetation fraction and photosynthesis through satellite remote sensing of visible and
82 infrared canopy reflectance (Kumar and Montieth, 1981). However, field observations

83 noted that canopy radiative transfer may not necessarily scale so neatly (Baldocchi *et*
84 *al*, 1985), leading to development of multiple canopy layer models (e.g., De Pury and
85 Farquhar, 1997). At the minimum, sunlit/shaded fractions of the canopy have to be treated
86 separately in models for accurate simulation of photosynthesis (Sprintsin *et al.*, 2012).
87 More sophisticated models now allow for multiple cohorts of interactively competing and
88 shading species with varying plant functional types (Medvigy *et al.*, 2009).

89 The primary production model in most ecosystem models now consists of a leaf-
90 level photosynthesis mechanism, embedded within a leaf boundary-layer coupling, leaf
91 energy balance model, canopy scaling algorithm, a soil water and humidity-sensitive
92 transpiration model, and sometimes a nutrient transformation and transport model,
93 primarily for nitrogen. Despite the apparent complete description of canopy
94 photosynthesis, interactions and small changes in parameters of these components causes
95 ecosystem models predict widely divergent estimates of the sensitivity of canopy
96 photosynthesis to climatic and biotic changes, even when they're using the same equations
97 (Schaefer *et al.*, 2012). Differences in parameters that controls rates of leaf respiration,
98 canopy architecture, or microclimate variation have large effects on canopy photosynthesis
99 rates and sensitivity.

100 Uncertainties also arise in our understanding of variation in radiation quality and
101 sun flecks, multiple interacting species, age-dependent changes in photosynthesis and
102 transpiration, moisture and nitrogen availability in soil, transformation of assimilated
103 carbon into storage pools, and canopy-scale stomatal and photosynthetic rate responses to
104 atmospheric CO₂ enrichment (e.g., Fig. 1). Models tend to underestimate variability in
105 canopy photosynthesis in response to climatic anomalies, but overestimate threshold

106 responses to climate and biospheric state shifts. On the one hand, ecosystems, as an
107 assemblage of species, and hence comprised of species that are adapted to have
108 compensating responses to environmental change are more conservative than the single
109 “model” species represented by the typical plant functional type approach in models. On
110 the other hand, overall ecosystem variance is large and since species present on the
111 landscape are partly a function of local climate and soils, they may show additive effects in
112 response to climate anomalies that exceed the a threshold (i.e., an extreme).

113 Consequently, there is a role for long-term canopy scale observations of
114 photosynthesis toward evaluating and improving these kinds of model responses. The rise
115 of canopy-scale observations of net ecosystem exchange of CO₂ (*NEE*) and inference of
116 gross primary productivity (*GPP*) through tower-based eddy covariance methods
117 (Baldocchi, 2008) and other canopy-scale experiments have expanded our ability to make
118 claims of model fidelity and sensitivity. For example, recent articles have focused on
119 dryness (Yi *et al.*, 2009), temperature (Niu *et al.*, 2012), light response (Schaefer *et al.*,
120 2012), phenology (Richardson *et al.*, 2012), short-term climate fluctuations (Medvigy *et al.*,
121 2010), interannual variations (Keenan *et al.*, 2012a), disturbance (Amiro *et al.*, 2010), and
122 CO₂ enrichment (Norby and Zak, 2011). Still, many of these studies are limited by short
123 time series at most sites. Advanced methods to link anomalies of environmental drivers to
124 fluxes have only been applied in limited cases.

125 In this manuscript, I focus on the value of longer-term observations of high-
126 frequency photosynthetic flux observations at the canopy scale. Are there effects of lagged
127 environmental drivers that are masked at shorter time scales? Can ecosystem adaptation
128 and community reorganization responses from extremes be detected? Hypotheses have

129 been put forward suggesting that canopy-scale productivity may be linked to prior season
130 carbon storage of non-structural or labile carbohydrates (Carbone *et al.*, 2007) and also
131 antecedent moisture availability in mesic forests (Ricciuto *et al.*, 2008; Desai *et al.*, 2010).
132 These lagged responses would not be seen at the leaf level, where direct controls of
133 moisture availability and leaf carbon content influence photosynthesis strongly by limiting
134 rates of light harvesting and total leaf area.

135 The purpose of the analysis here is not to put forth brand new models of canopy
136 photosynthesis, but rather to highlight the path forward with long-term observations. As
137 such, my goals are twofold: 1) to demonstrate advanced statistical methods to evaluate
138 modes of variation of long-term environmental data and 2) observe how these methods
139 provide new insight into lags and switches of canopy photosynthesis that make it so hard
140 to model and so different from leaf-level responses. In particular, I use a 15-year record of
141 regional NEE from a very tall tower in the north central US to test which of antecedent soil
142 moisture availability and prior productivity most influence canopy productivity. Moisture
143 lags likely represent the accumulation and storage of available water that allows a canopy
144 to respond to periods of high water demand or physiological stress. Similarly, prior carbon
145 accumulation influences current carbon uptake both in changes in allocation and
146 development of carbohydrate reserves. Can these effects be seen in observations and if so,
147 which is the most relevant to include in models?

148 I hypothesize that short-term (daily) drivers of photosynthesis are primarily light
149 and temperature, but longer-term (weekly to annual) is limited primarily by moisture and
150 internal storage of prior photosynthate. These hypotheses are tested against the landscape-
151 scale observations of net carbon uptake and associated surface and meteorological forcing.

152 **Methods**

153 **Analysis Framework**

154 Many of the analyses described in the prior section attempted to use many short-
155 term (2-5 year) ecosystem flux observations to improve prediction or modeling by
156 substituting space for time (e.g., Yi *et al.*, 2010). Many sites each with a few years of data
157 are analyzed to make inferences about long-term evolution of biosphere to climate and
158 environmental drivers. However, there is some evidence that predictions made in this form
159 across sites do not necessarily map well onto long-term predictions at a single site (Desai,
160 2010; Keenan *et al.*, 2012a). The limiting factor is the lack of long-term high-frequency
161 observation of the state of the biosphere.

162 Short-term measurements are difficult to use for diagnosing anomalies and
163 extremes. Short-term multi-site studies can diagnose mean state and mean variability of
164 carbon fluxes, but may underestimate the true level of variability over years and how
165 extremes in climate and biotic disturbance (including both large short-term pulse and long-
166 term press (steady pressure) disturbances) can drive carbon assimilation differently than
167 short-term responses. For example, many eddy covariance flux tower studies focus on a
168 few years of data to identify particular climate responses (e.g., wet year versus dry year),
169 but are likely to have confounded co-variability among climate factors, to miss multi-year
170 responses, or and have low probability of capturing extreme climate events and the role of
171 pre-conditioning of ecosystem states. Community-scale response from changes in resource
172 availability and competitive advantages to these can occur in forest ecosystems at
173 timescales of years to decades (Gellesch *et al.*, 2013).

174 While some processes like photosynthetic acclimation have been well captured in
175 lab experiments, they are harder to diagnose with short-term environmental data, even
176 with a decade long record from enriched CO₂ experiments (Norby and Zak, 2011). The
177 statistical sample for environmental observations of low probability, high impact events is
178 too small.

179 A benefit of an evolving measurement network is that over time there are sites that
180 start having long records where one can look closely at features like memory effects (long
181 lag relationships), decadal trends, and state shifts that would not be easily noted across
182 space. Instead, the challenge is addressing the data deluge. A decade long flux tower record
183 of half-hourly NEE observations and related climate drivers can easily exceed 10⁶
184 observations. Moreover, the data are strongly auto-correlated and may suffer from
185 harmonization issues related to changes in instruments, measurement height, and so forth.
186 As a consequence, many analyses, even at long-term tower sites, limit their analyses to
187 subsets of the data. For example, Keenan *et al* (2012b) found no suitable combination of
188 parameters of a simple model could adequately explain three separate five-year periods in
189 NEE observed over the 18-year record at Harvard forest.

190 There is ongoing work on improving harmonization of long-term datasets like
191 decadal eddy covariance and the evolving National Ecological Observatory Network
192 (NEON), which will include nearly 60 sites across North America with eventually 30+ years
193 of carbon cycle and biological observations. These observations provide a suitable data
194 testbed if and only if the community first develops reliable and usable statistical metrics
195 and model-data evaluation. Therefore, in this study, I specifically focus on the more than

196 15-year record of eddy covariance carbon and water regional flux observations at a
197 forested site in the north central US (Figs. 2 and 3).

198 **Site Description and Data**

199 I analyzed 15-years of flux tower observations of CO₂ and H₂O flux from one of the
200 longest continuously running eddy covariance flux towers in the U.S., the WLEF Park Falls
201 tower (US-PFa) (Davis *et al.*, 2003), where fluxes have been measured since late 1996 with
202 minimal disruptions, except in 2002 (Fig. 2). Meteorological variables were also observed
203 at the site (Fig. 3 and Table 1). WLEF is unique for being the tallest flux tower across the
204 Fluxnet network, allowing us to observe the impact of patchy landscapes and canopy
205 interactions on carbon assimilation. My collaborators and I have observed fluxes at three
206 heights (30 m, 122m, 396 m) and use these to develop a single “preferred” flux product
207 (Davis *et al.*, 2003), based on boundary layer turbulence conditions. The tower samples a
208 fetch on the order of 1-5 km depending on atmospheric stability and wind speed.

209 Unlike canopy-scale towers, tall-towers sample fluxes that represent many species
210 and many soil types. However, an advantage of these observations is they are at a similar
211 scale to that which ecosystem models represent canopies and plant functional types (10s to
212 100s of km). Schafer *et al.* (2012) noted that ecosystem model estimates of daily GPP were
213 surprisingly well simulated at this site in a large flux tower-model intercomparison of GPP,
214 either because modelers have used this site significantly for calibration, or, that the fluxes
215 better represent the “model organism” being represented by the single plant functional
216 types used in most models.

217 The site samples carbon and water fluxes from a temperate mixed forest landscape
218 that consists of approximately $\frac{3}{4}$ forest equal parts young to intermediate age

219 commercially harvested aspen, mature northern hardwood (sugar maple, ash, basswood),
220 and red pine plantations (Desai *et al.*, 2007). The remaining $\frac{1}{4}$ is primarily a mosaic of
221 wetlands and shrub areas, including black spruce and peat bogs, cedar swamps, sedge
222 wetlands, and shrub fens. Spatial variability occurs in relatively small scales, driven by
223 microtopography and land management, while the overall landscape topography is flat and
224 density of human settlement in the tower footprint is minimal.

225 **Estimating Canopy-Scale Photosynthesis**

226 Eddy covariance towers observe the *net* exchange of trace gases, heat, and
227 momentum from the surface to atmosphere, based on well-established micrometeorological
228 theory (Baldocchi, 2008). Turbulence properties of the atmospheric surface layer allows
229 one to take the 30-60 minute mean covariance of high-frequency (>10 Hz) observations of
230 vertical wind and the flux tracer of interest (e.g., carbon dioxide, water, temperature)
231 summed with below-sensor net tracer storage and vertical flux divergence to represent the
232 net surface flux. Sonic anemometry (measuring vertical and horizontal wind components
233 with sound pulses) and infrared gas analyzers sampling air near the anemometer are
234 typically used to measure this net covariance. Contributions from low-frequency transport
235 (advection) are usually neglected, but tend to be small, of the same magnitude as the 10-
236 20% inherent random flux error (Yi *et al.*, 2000). Over the years, researchers have
237 instrumented nearly 500 of these sites for carbon and water cycle observations and general
238 quality control approaches have been identified for instrument noise, lag, and spectral
239 corrections, coordinate geometry rotation for wind velocity, low turbulence screening, and
240 other turbulence statistics, which are applied here (Berger *et al.*, 2001; Foken *et al.*, 2012).

241 The focus of this analysis of the effect of climate anomalies on photosynthesis, not
242 net exchange (which includes respiration and decomposition processes). Therefore, I
243 developed a method to represent this photosynthesis from net ecosystem exchange of CO₂
244 (NEE). Unfortunately, there is no single accepted method for doing so, and all require some
245 level of empirical assumptions or statistical inference that partly takes advantage of the
246 lack of GPP at night. Consequently, methods diverge on estimates of GPP by more than 20%
247 and can include artifacts from fitting NEE to respiration models (Desai *et al.*, 2008).

248 Since I want to focus on the value of NEE to models, I developed an alternate metric
249 of the effect of canopy photosynthesis on NEE, termed net photosynthetic drawdown (P_d), a
250 daily metric of canopy photosynthesis that removes assumptions used in many GPP models.
251 P_d was estimated at a daily timescale from the hourly flux data as the difference in
252 nighttime to daytime NEE. Maximum nighttime NEE was identified at night when more
253 than four hours of good observations were available. Maximum is used over mean since it
254 has been shown to be closer to the advection corrected observations at night (Van Gorsel *et*
255 *al.*, 2009). This estimate of nighttime NEE is then differenced with the mean daytime gap-
256 filled NEE between 10 and 14 local time if there are more than four hours of good
257 observations during that day (when the sun is up). Here I use gap-filled NEE to avoid
258 biasing the mean NEE, which exhibits a strong diurnal cycle. Gap-filling errors tend to be
259 much smaller than GPP uncertainty (Moffat *et al.*, 2007). The P_d time series is shown in Fig.
260 4a. The P_d time series has 5,490 days of data, with 37% of data missing.

261 Further analysis showed that use of P_d instead of GPP does not significantly change
262 the results or conclusions of this study and presents a novel way to understand the effect of
263 climate on photosynthesis. The correlation of P_d to GPP is high, particularly for maximum

264 daily GPP ($r^2=0.81$) and greater at the monthly timescale ($r^2=0.96$). The fit is linear for GPP,
265 with an intercept of 0 (Fig. 5). Since P_d is a detector of maximum daily photosynthetic
266 uptake and has a greater dynamic range than GPP, it is likely that P_d is better at detecting
267 extreme photosynthesis responses to climate anomalies. While this method is conceptually
268 analogue to atmospheric CO₂ “drawdown” (e.g., Desai *et al.*, 2010), it is different as the flux
269 drawdown does not include covariation with boundary layer depth and represents a much
270 smaller footprint.

271 **Statistical Analysis**

272 I tested the hypotheses mentioned above by testing for both direct and lagged
273 relationships between P_d and climate forcing factors (Table 1) at multiple time scales and
274 compared them to the autocorrelation of P_d . A number of studies have identified
275 characteristic timescales of variability in flux data using wavelet, single spectrum, or
276 Fourier time-series analysis (e.g., Baldocchi *et al.*, 2001; Mahecha *et al.*, 2007; Sevanta and
277 Williams, 2009; Stoy *et al.*, 2009), which have all noted characteristic peaks of variability in
278 NEE especially at the diurnal, synoptic (3-4 day), seasonal, and interannual timescale.
279 Similarly, frequency dependent model-data comparisons (e.g., Dietze *et al.*, 2012; Mahecha
280 *et al.*, 2010; Keenan *et al.*, 2012a) have all found deficiencies of models in representing
281 many of these modes of variability.

282 I identified these scales in daily P_d and evapotranspiration (ET) flux using a similar
283 analysis of empirical model decomposition (EMD), whose results are fed into the Hilbert-
284 Huang spectral transformation (HHT) (Huang and Wu, 2008). EMD is an empirical
285 approach to time series deconvolution that does not require assumptions of cyclical
286 behavior (as needed by Fourier) or stationarity and does not require determination of the

287 shape of the weighting kernel or wavelet. The discontinuous EMD (Barnhart *et al.*, 2012)
288 further extends the application to time series with missing data by applying a mirroring
289 approach to fill the data gaps. EMD decomposes a time series into a series of intrinsic mode
290 functions (IMF) also in the time dimension, which when fed to the HHT algorithm that
291 outputs a time by frequency power spectrum.

292 Timescales for analysis were determined from the HHT of P_d and ET (Fig. 4). Both
293 signals have a number of similar modes of variability, especially at the synoptic, monthly,
294 and seasonal timescale. ET has greater temporal variations in these modes and greater
295 signal on long-time scales (> 100 days). Longer timescale variability is present in the
296 growing season more than outside of it. HHT identified strong monthly peaks that were not
297 previously identified and suggests that interannual variability explains less of the signal
298 than other methods have previously shown (e.g., Baldocchi *et al.*, 2001). Methodologically, I
299 used this analysis select averaging timescales of 1, 3, 8, 15, 30, 90, 180, 360, 720, and 1440
300 day, as described next.

301 The HHT analysis also identified the importance of normalizing variability across
302 timescale to best identify climatic and internal controls on P_d . For example, there is a
303 variety of literature that will show high correlation of GPP to other variables, solely
304 because the main modes of variability (e.g., the annual solar cycle) are strong in both, not
305 because one truly explains the other. This method of analysis is disingenuous when it
306 comes to the question I seek to answer here.

307 One of the benefits of long-term data is the ability to remove much of this co-
308 variability and look at how anomalies or extremes manifest themselves in the data and how
309 they are correlated to anomalies or extremes in another variable. If a daily time series

310 signal $X(day, year)$ is stationary (as it mostly appears to be in this case and discussed more
 311 in the discussion), then an anomaly time series $X_{an}(day, year)$ can be simply constructed by
 312 removing the ensemble mean:

$$313 \quad X_{an}(day, year) = X(day, year) - \overline{X(day)} \Big|_{year} \quad (1)$$

314 where $X(day)/Y$ is the daily time series of variable X ensemble averaged across all years. To
 315 test across multiple timescales, I applied a forward averaging filter across the time series
 316 (e.g., Fig. 6b), avoiding forecasting by removing the end of the data series:

$$317 \quad X_{an}(day, year) \Big|_{timescale} = \frac{1}{timescale} \sum_{t=day}^{t=day+timescale} X_{an}(t, year) \quad (2)$$

318 where timescale is the number of days to average. One issue that arises when analyzing this
 319 variable across seasonal to interannual timescales is the need for averages to stay aligned
 320 with the solar orbital forcing cycle, so that averages in any one year can comparable to
 321 other years. To do this, I reduced each year to 360 day length by removing the first few and
 322 last few days of data for each year and choosing averaging scales which share divisors with
 323 360. The choice of timing has relatively minimal effect and the choice of winter, where
 324 carbon fluxes are near zero is ideal. A second issue involves gaps in the data. For gaps, I
 325 sampled the data with replacement, filling gaps linearly across small gaps (days) and taking
 326 long-term means for longer gaps (weeks).

327 To remove the previously discussed solar forcing driven artificial correlation among
 328 variables, I normalized the time series. Flux anomalies (e.g., Fig. 6c) display a strong
 329 seasonality given the change in variability from winter to summer (Fig. 3). Relative
 330 anomalies $X_{rel,an}(day, year) \Big|_{timescale}$ were derived from averaged anomalies by dividing the

331 time series by the ensemble average standard deviation across all years for a given
332 averaging timescale:

$$333 \quad X_{rel,an}(day, year) \Big|_{timescale} = \frac{X_{an}(day, year) \Big|_{timescale}}{\left(X_{an}(day, year) \Big|_{timescale} \right) \Big|_{year}} \quad (3)$$

334 The remaining time series appears stationary and random (Fig. 6c), and reflects a
335 statistically defensible view of relative anomalies of the time series as a function of
336 averaging filter. Relative anomalies in this fashion were computed for P_d (Fig. 6) and a
337 variety of observations to test hypotheses including variables related to canopy physiology
338 and structure such as remotely sensed vegetation index (EVI) and minimum, maximum,
339 mean, and diurnal range of air temperature, remotely sensed land surface temperature
340 (LST), and variables related to canopy moisture availability including ET , water use
341 efficiency (WUE , GPP divided by ET), precipitation, and soil moisture, as noted in Table 1
342 and shown in Fig. 3. Remotely sensed variables were derived from the NASA MODIS TERRA
343 and AQUA reflectance properties and downloaded from the ORNL MODIS land product
344 subset server (<http://daac.ornl.gov/MODIS/>), while other variables were directly observed
345 by the tower with gaps filled from harmonized daily climate data downloaded from the
346 National Climatic Data Center archive of National Weather Service co-operative observer
347 stations and weather forecast reanalysis from the NOAA North American Regional
348 Reanalysis (NARR).

349 I compared relative anomalies of all variables to relative anomalies of P_d at all
350 averaging timescales both with direct linear correlation and with lagged correlation where
351 climate factors were lagged against P_d at a range of lags equivalent to the averaging

352 timescales. A two-tailed t-test for significance was applied to all correlation coefficients and
 353 only those coefficients that were significant at the 90% level were saved.

354 The significance test was modified to account for the autocorrelation present in all
 355 environmental time series. Consequently, the degrees of freedom to apply to significance
 356 tests should be much smaller than the total number of samples. I reduced the degrees of
 357 freedom using a modified effective degrees of freedom (EDOF) approach of Bretherton *et al.*
 358 (1999):

$$359 \quad N_* = \frac{N}{\sum_{t=N/2}^N \left[\left(1 - \frac{t}{N} \right) \rho_t^X \rho_t^Y \right]} \quad (4)$$

360 where N_* is the reduced degrees of freedom for significance testing of correlation of two
 361 time series X and Y with N samples. ρ_t^X represents the autocorrelation of time series X at lag
 362 t . Though most of the autocorrelation is in the first few lags, I included all lags to $N/2$ to
 363 account for long lead correlations. Further, only those correlations of variables to P_d that
 364 exceeded the lagged autocorrelation of P_d are used as a test to compare moisture versus
 365 carbon storage control as predictors of current P_d .

366 The EDOF of P_d (Fig. 7) reveals that while daily P_d has over 3400 observations, the
 367 EDOF at daily scale is only slightly above 600, and decreases nearly linearly with
 368 logarithmic increases in averaging timescale, such that interannual analysis is limited to
 369 EDOF in the few tens to single digits. As shown in the results, this limits the ability of this
 370 analysis to diagnose correlations of anomalies at multi-year timescales and highlights how
 371 even a 15 year time series may be unreliable for detecting interannual and longer trends
 372 and correlations. The results here shows that a few time series were able to meet

373 significance threshold, but multi-year correlation analysis with flux tower data requires
374 sufficiently long data sets because of high auto-correlation.

375 A second test was also applied for comparing predictive ability of variables against
376 inherent autocorrelation of P_d , known as the Granger causality analysis (Detto *et al.*, 2012).
377 The method originated from economics, but has recently gained popularity in geophysical
378 time series analysis. The analysis builds a multiple linear regression of lagged values of a
379 time series X to predict current values of X . This regression is then compared iteratively to
380 including an increasing number of lagged values of time series Y that are significantly (two
381 tailed t-test at 90% level) correlated to X to predict current values of X . When the new
382 regression significantly improves upon (tested with an F-test) the autocorrelation
383 regression, those lags of Y are retained and the terminology is that Y “Granger causes” X , or
384 prior values of X and Y at certain lags explains a significantly larger fraction of X than the
385 prior values of X does by itself. A minor modification was made here to include the reduced
386 degrees of freedom for significance testing and the replacement of the F-test with the more
387 empirical Aikake Information Criterion, which incorporates both the likelihood of the
388 regression and penalties for number of parameters. Here, I tested Granger causality for all
389 variables against P_d (Table 2).

390 Though all of these methods are relatively standard time series analysis, they have
391 only recently been applied to carbon flux data, mostly because long time series of flux data
392 are only now becoming common. The results show how the lagged correlation, spectral,
393 and causality analysis together provide insight on how canopy photosynthesis is different
394 from leaf photosynthesis and how it can be leveraged to improve canopy photosynthesis
395 models.

396 Results

397 Modes of variability in observations

398 Large variability exists in hourly flux data of *NEE* and *ET* (Fig. 2). Outliers exist in
399 most years, but positive anomalies in *ET* and *NEE* decreased after 2005. Diurnal and
400 seasonal variability dominate the signal and trends. The decrease in *NEE* uptake from
401 2006-2010 is visually evident, but difficult to discern quantitatively against the variability
402 at the hourly scale (Desai *et al.*, 2010). Not surprising, hourly data can be challenging when
403 it comes to statistical data assimilation approaches to constraining ecosystem models
404 (Zobitz *et al.*, 2011), as these methods tend to force model parameters toward the
405 dominant modes of variability (diurnal and seasonal) and limit excursions away from the
406 mean (over-fitting the mean at the expense of the extremes) (Desai, 2010) especially when
407 data uncertainty is large, as it is the case for hourly flux observations (Raupach *et al.*, 2005).
408 Not surprisingly, models have great difficulty simulating other temporal modes of
409 variability in photosynthesis (Dietze *et al.*, 2012; Keenan *et al.*, 2012a). Moreover, many of
410 the anomalies of photosynthesis that are related to climate anomalies may exist only at
411 longer timescales.

412 The EMD analysis directly shows the importance of weekly to seasonal variability in
413 *P_d* and *ET* (Fig. 3), counter to previous wavelet based analyses. For example, Baldocchi *et al.*
414 (2001) found a spectral gap in flux tower *NEE* at the three-four week scale. The
415 transformed data and the more empirical approach of HHT reveal that there are variations
416 present at this scale, perhaps more so than at longer scales, and not consistently at a fixed
417 value. Longer-term variability (90-360 day) is more present in the growing season

418 (summer) and more distributed across frequency for ET compared to P_d , suggesting more
419 coherence in interannual variability of ET than P_d .

420 **Direct Relationships**

421 Given the range of variability in flux tower observations and a general overlap of
422 modes of variability in ET and P_d anomalies at the daily to seasonal timescales, it should be
423 expected that some level of correlation exists among these factors. The strongest direct
424 linear significant correlations between ET and P_d anomalies exist at scales between 30-360
425 days, but persist down to 1 day, where ET explains around 10% of the variability in P_d at
426 short timescales and approaches 40% at seasonal scales (Fig. 8). Longer-term correlations
427 are not significant, but this may partly be the result of insufficient length of time series and
428 the strict degree of freedom constraint. A similar level correlation is seen for WUE , but not
429 at annual time scales. This correlation partly stems from self-correlation since GPP , which
430 goes into the WUE calculation, and P_d are both derived from NEE .

431 Temperature is a well-known factor influencing leaf-level photosynthesis. Even
432 though solar forcing leads to a strong correlation between temperature and NEE or GPP at
433 diurnal and seasonal scales, when the ensemble average is removed and the variables
434 compared as standardized anomalies, the relationship is much weaker, though it remains
435 significant from 1 day out to 30 day, with r around 0.15. None of the temperature factors
436 (average, maximum, minimum, mean, range, or land surface temperature) are particularly
437 better than others at explaining variation in P_d , though stronger correlations exist for
438 maximum daily temperature especially at longer timescales, including a particular strong
439 correlation at the 1440 day scale. It should be noted that LST is provided at 8-day intervals,

440 so 1 and 3 day average correlations are not included. For the purpose of the lag analysis,
441 given similar relationships, only T_{mean} is further analyzed.

442 For direct correlation, the results are weaker for soil moisture than for temperature,
443 but do provide some evidence to the importance of seasonal moisture budgets on net
444 carbon assimilation, even in mesic forest/wetland systems. This result less consistent with
445 an earlier model calibration study that showed interannual variability in NEE at WLEF was
446 best explained by soil moisture (Ricciuto *et al.*, 2008), but when lags are taken into account,
447 the results change, as shown below. Though not shown, results with atmospheric humidity
448 variables were similar to those of Q_{soil} . Precipitation anomalies only weakly P_d at 3-day
449 averages. This result may be caused by the intermittent nature of precipitation, the greater
450 error in short term precipitation, and the importance of soil percolation processes prior to
451 plant water uptake. Finally, EVI anomalies are found to have no correlation to P_d anomalies
452 at any timescale.

453 Lagged analysis

454 Lagged analysis (Fig. 9 and Fig. 10) further supports results shown in Fig. 8, but also
455 reveals subtleties regarding moisture. The non-moisture specific variables support both a
456 strong autocorrelation of P_d as previously shown in Fig. 6 and lack of correlation at many
457 timescales for EVI anomalies. Lagged EVI (Fig. 9b) does have some weak, but significant
458 negative correlation to P_d for monthly averaging timescales and three month lag, and this
459 correlation exceeds the autocorrelation. This signal represents both the impact of summer
460 vegetation stress (low EVI) on autumn photosynthesis (reduced) (Wu *et al.*, in press) and
461 the impact of phenology (late spring = low EVI) on net carbon uptake (reduced). Spring
462 flush is typically a two-week process and the growing season is around three months.

463 The autocorrelation of P_d (Fig. 9a) is persistent at AR-1 and AR-2 out to 30 days. For
464 any given lag, some amount of autocorrelation exists for all averaging times up to the lag,
465 increasing as the lag approaches the averaging timescale (i.e., AR-1). Some out of phase
466 (180 day) negative correlation exists as it also does at the very long timescale (1440 day),
467 the latter of which is difficult to explain. At one level, ecosystem models incorporate this
468 autocorrelation through the “memory effect” of labile carbon pools, but these are probably
469 not responsible for the observed longer timescale correlations, possibly tied to other non-
470 structural carbohydrate pools (e.g., Carbone *et al.*, 2007) or signals of community
471 reorganization, ecosystem dynamics, or climate oscillations. However, the lack of long
472 autocorrelations in P_d do not provide much evidence here for strong long-term internal
473 control through non-structural carbohydrates, shifts in plant allocation, or community
474 reorganization.

475 Lagged temperature is predictive for P_d anomalies at short lags and averaging times
476 out to a few weeks (Fig. 9c and d), with a positive correlation (warm anomalies lead to
477 increased P_d one to two weeks out), but these are much weaker than the autocorrelation.
478 An interesting weak, but stronger than autocorrelation, relationship exists with 90-day
479 prior daily mean temperature and current P_d , hinting at possible long-lag effects, where a
480 short early season warm spell (e.g., false spring), can enhance growth. There is also a weak
481 but significant negative relationship of seasonal average temperature from four years past,
482 once again difficult to explain.

483 Moisture anomalies (ET , WUE , P_{recip} , and Q_{soil}) also have some predicative ability for
484 P_d anomalies and most all positive – increased ET or precipitation enhances carbon
485 assimilation. Correlations of lagged ET to P_d (Fig. 9a) exist at short timescales and the first

486 two lags for each timescale, but they are surprisingly weak compared to the strong direct
487 correlation (Fig. 8). Similar to temperature, a four-year lag negative correlation of ET on P_d
488 is found for seasonal averaging.

489 Relationships also exist in weekly to seasonal average precipitation and soil
490 moisture at the 2-3 month lag. Positive anomalies in soil moisture are more predictive for
491 future weekly to seasonal P_d than P_d itself, suggesting a long-term moisture control. For
492 example, early season weekly to seasonal moisture deficits inhibit end of season carbon
493 assimilation. This effect is of slightly greater magnitude and correlation as the effect of
494 direct moisture deficits on P_d . Unlike the direct correlation, these results more strongly
495 support both the work of Ricciuto *et al.* (2008) and the second hypothesis of long-term
496 moisture control on P_d .

497 Causality analysis

498 The findings of the Granger causality analysis (Table 2) are consistent with the lag
499 analyses for most except the moisture variables. Daily to weekly temperature and ET both
500 Granger cause P_d , with additional longer term control by seasonal EVI and annual ET.
501 Interestingly, Q_{soil} does not Granger cause P_d at any timescale. Only ET anomalies have
502 predictive ability at long time scales and highlights the difficulty of dissecting the causes of
503 seasonal to interannual variability even with 15 years of flux tower data.

504 Discussion

505 Role of lagged forcing in photosynthesis

506 Internal control (carbon assimilation rates related to prior carbon assimilation
507 rates) is a strong predictor on canopy carbon assimilation at timescales up to a week, while
508 the key direct climatic modifier of this are temperature and available moisture, but

509 primarily on longer timescales and longer lag times. Clearly, this differs from the primarily
510 direct relationships one finds for leaf photosynthesis controls and highlights the
511 complexity of modeling canopy photosynthesis and the value of long-term data.

512 The analysis here is unable to directly identify mechanisms. I assumed that a strong
513 autocorrelation at long time scales implies carbohydrate storage or other mechanisms of
514 buffering that limits ecosystem response to climate extremes. However, the evidence that
515 such occurs here is weak and instead long-lead short-term moisture stress and prior
516 season temperature anomalies appear to have the strongest effects, suggesting that
517 moisture control is stronger than expected. A number of biotic interactions are likely in
518 response to climate anomalies, in addition to changes in internal storage of sugars and
519 starches, there are possible shifts in allocation in response to extremes or aging or
520 community reorganization from shifts in competitive pressure. There was very limited
521 management in the region over the study period and no evidence for a shift in age structure
522 or dominance of certain ecosystems (Gellesch et al., 2013; Scheffer et al., 2001), which
523 allowed this analysis to assume stationary conditions. Additionally, it is unclear from the
524 analysis if any state shifts from multiple or repeating stress, dubbed “ecological stress
525 memory” (Walter et al., 2013), was observed here. Instead, the analysis suggests that
526 modeling and experimental studies should look and evaluate carbon cycle shifts over long
527 time scales (seasons to years) in response to relatively short-term drought manipulation or
528 prior season temperature anomalies. Further, multi-year anomalies that may be related to
529 patterns and oscillations in biology or climate warrant more investigation. For example, the
530 negative autocorrelation of P_d at interannual timescales for annual and longer averages

531 suggests potential biological oscillations, cycles of herbivory, species successional
532 processes.

533 Further evidence on the lack of a strong negative feedback or “internal control” is
534 the surprising lack of correlation between relative anomalies of P_d to EVI . Anomalies of P_d
535 do not appear to relate to anomalies of EVI at any timescale, calling into question how well
536 remote sensing can be used to evaluate how climate anomalies drive productivity
537 anomalies. Many applications have been developed around the ability to apply differences
538 in infrared and visible reflectance of canopies to estimate global photosynthesis from space,
539 ever since early work showed the strong link of absorbed radiation to plant carbon
540 assimilation (e.g., Kumar and Monteith, 1981). For example, remarkably strong monthly to
541 seasonal correlations exist between NASA MODIS derived monthly GPP against flux tower
542 estimated GPP (Heinsch *et al.*, 2006). However, many of these papers find that while
543 satellites can sense large-scale latitudinal variation, significant unexplained variability
544 exists across smaller regions and across longer timescales.

545 It appears that EVI and similar metrics of remotely sensed vegetation greenness or
546 absorbed radiation capture processes like phenology, leaf area, or canopy development,
547 though they likely do not readily capture the anomalies or extremes as formulated in this
548 analysis. Though anomalies do not correlate, EVI does explain approximately 75% of the
549 biweekly variation of P_d . EVI has also been argued as a good proxy for carbon uptake
550 phenology, but at this site, dates of start and end of carbon uptake period (period when
551 mean daily smoother P_d is positive) do not correlate strongly to dates of start and end of
552 the “greenness” period as identified in EVI . However, I did find that growing season length
553 as defined by its carbon uptake period has a strong correlation with average growing

554 season EVI ($r=-0.88$), though with a negative relationship and a small effect size, suggesting
555 that short term and long-term EVI have opposing effects and may partially explain the lack
556 of correlation of anomalies. The strongest relationship with mean growing season EVI is
557 with correlation of the end date of this uptake period ($r=-0.92$), consistent with some
558 recent work that many temperate forest systems have interannual variability in NEE driven
559 by end of season signals (Wu *et al.*, in press). The analysis here does suggest caution is
560 warranted when analyzing anomalies in broadband satellite vegetation indices.

561 Finally, it is apparent that both at the direct timescale, through the high correlation
562 of ET to P_d and at the lagged timescale, through the positive association of seasonal soil
563 moisture to P_d and the long lag relationships of temperature to P_d , all imply a variety of
564 moisture retention and moisture use processes influence photosynthesis at a number of
565 timescales. Some of these maybe related to summer droughts influence late season
566 photosynthesis and others may be related the dynamics of the snowpack on soil moisture.
567 The existence of moisture control on plant biogeochemistry in a mesic temperate forest
568 and wetland landscape in and of itself is surprising and opens up a number of new avenues
569 for analysis.

570 **Towards a canopy photosynthesis modeling framework**

571 Compared to lab experiments, both uncontrolled and controlled field observations
572 require greater explicit consideration of time and spatial scale, and the extent to which
573 variability expressed in one dimension truly reflects the signal one seeks to estimate. This
574 paper, like others (e.g., Stoy *et al.*, 2009), demonstrate that frequency dependent analysis is
575 essential for identifying processes over long-time periods or large regions. Otherwise,

576 conclusions can be drawn from short-term or small-scale data that have very limited
577 application to how photosynthesis actually responds to the environment.

578 Statistical analyses for large environmental data sets are still in development.
579 Modern computational speeds, open source libraries for advanced programming languages,
580 and new models of graduate student training have led to continued improvement in these
581 (Zobitz *et al.*, 2011). Of course, whether the methods presented here are useful ultimately
582 depend on the interpretation of results.

583 Moorcroft *et al* (2006) asked if we have reached a predictive ability for the
584 biosphere. Progress has been made, especially with advanced coupled dynamic vegetation
585 and carbon cycling models (e.g., Medvigy *et al.*, 2009), but the community may have
586 reached a standstill until we seriously reconsider how we confront models with data.
587 Dietze *et al* (2012) found among more than a dozen ecosystem models, very little
588 confidence present in our ability to simulate both the diurnal cycle and interannual
589 variability, with the latter finding confirmed by Keenan *et al* (2012). New research further
590 finds even short-term environmental variability can strongly affect long-term carbon
591 cycling (Medvigy *et al.*, 2010).

592 The analysis here suggests that models need to be evaluated on the temporal
593 memory of moisture and carbon storage mechanisms. Advances have been made in
594 applying data assimilation or Bayesian inference methods to sift through data and models
595 (Williams *et al.*, 2009). Large model-data syntheses as mentioned in the introduction have
596 contributed to our ability to diagnose consistent model errors. Data uncertainty, machine
597 readability, and archival have also gotten greater attention. Uncertainty, in particular, is
598 essential to collect with all these data, given how sensitive model-data comparisons can be

599 to them (Raupach *et al.*, 2005). Finally, recent progress has been made on making modeling
600 and model-data comparisons a routine exercise, or at least, more user-friendly and across a
601 wider range of data sources (LeBauer *et al.*, in press).

602 Multiple lag and time filtering should applied to climate extreme experiments
603 conducted with ecosystem models in the soil moisture and carbon storage domains.
604 Various model structures and parameters (e.g., root exudates, labile carbohydrate storage,
605 community shifts, or soil moisture storage rates) can be investigated not for merely how
606 well they simulate *NEE* or even *P_a*, but rather how well they simulate the observed
607 relationships among variables across time. Some of these responses span over multiple
608 years. Further, comparisons should be made in anomaly space if we really want to test how
609 climate extremes influence photosynthesis.

610 **Conclusion**

611 I found that neither the carbon storage control or moisture control hypotheses
612 could be falsified with long-term data, once seasonal cycle was removed. The extent to
613 which the hypotheses could be falsified was strongly dependent on the scale of the
614 averaging filter and the lags analyzed. Using Hilbert spectra to identify relevant lags, I
615 found a short-term carbon storage link on the order of weeks and a longer-term seasonal
616 positive soil moisture influence on photosynthesis anomalies. Daily to weekly lagged
617 positive anomalies of photosynthesis positively influence current photosynthesis,
618 inhibiting photosynthetic response to direct climate extremes, primarily anomalies on
619 evapotranspiration and maximum temperature. Moisture stress or surplus in the prior
620 season did inhibit or promote photosynthesis, but mechanisms are difficult to assess.

621 These results support prior suppositions that spring moisture anomalies and
622 autumn carbon uptake anomalies influence future carbon assimilation rates, not just length
623 of growing season or phenology. Further, the results highlight the difficulty that some
624 commonly used indicators of plant growth such as remotely sensed vegetation indices, can
625 reliably detecting anomalies in net carbon uptake. Finally, multi-year lagged negative
626 relationships of temperature and evapotranspiration anomalies on current photosynthesis
627 are intriguing and suggest new avenues of exploration for the role of long-lead ecosystem
628 responses to extremes.

629 These findings are not necessarily detectable with shorter-term data or leaf-level
630 analysis, as they involve subtle relationships and canopy and soil level processes. The
631 results are similar to, for example, recent work by Niu *et al.* (2012) who argued that
632 thermal acclimation of NEE (a flux made up of many interacting processes) can occur on
633 interannual timescales in canopies. Also interesting was the lack of relationship between
634 spectral indices and P_d , once converted to anomaly space, similar to some of the results of
635 Heinsch *et al* (2006) that remotely sensed *GPP* is adequate for large spatial scale variation
636 but poor for single pixel interannual variability.

637 The results here demonstrate the importance of long-term environmental
638 observation of canopy photosynthesis but the caution that has to be taken regarding the
639 high temporal autocorrelation that exists in flux and climate data. Strong covariance of
640 these signals to seasonal orbital forcing requires careful evaluation of spurious correlation.
641 A disconcerting finding was the lack of strong statistical power at detecting many
642 correlations at long timescales, beyond interannual, even with >15 years of data. Methods
643 that seek complementary use of short-term field manipulations, lab observations, and long-

644 term datasets like Fluxnet and the evolving NEON observatory will require continued
645 evolution of model-data comparison tools. Other papers in this special issue point to a
646 number of intriguing new ways to look at photosynthesis in models (Dietze *et al.*, this issue;
647 Rogers *et al.*, this issue; Sitch *et al.*, this issue; Tholen *et al.*, this issue). Here, I have shown
648 that a spectral anomaly framework and long-term flux observation network contribute to
649 their evaluation and improvement.

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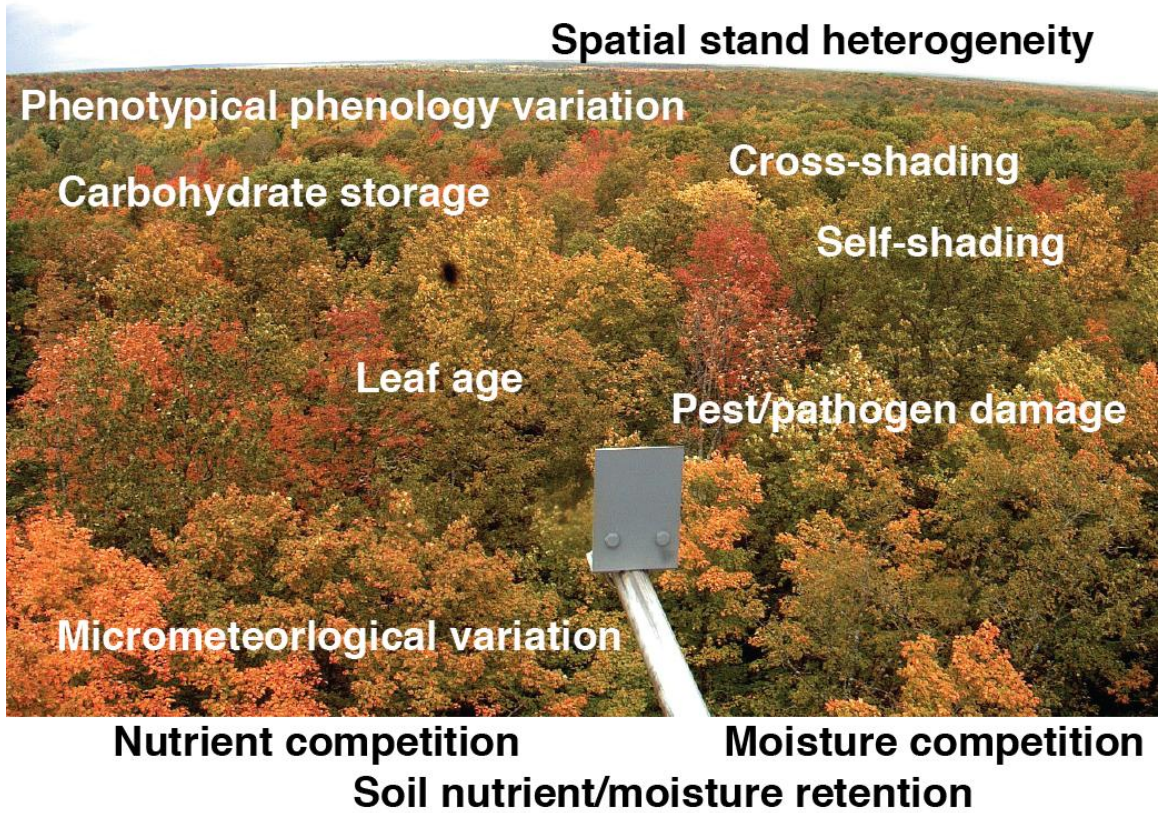
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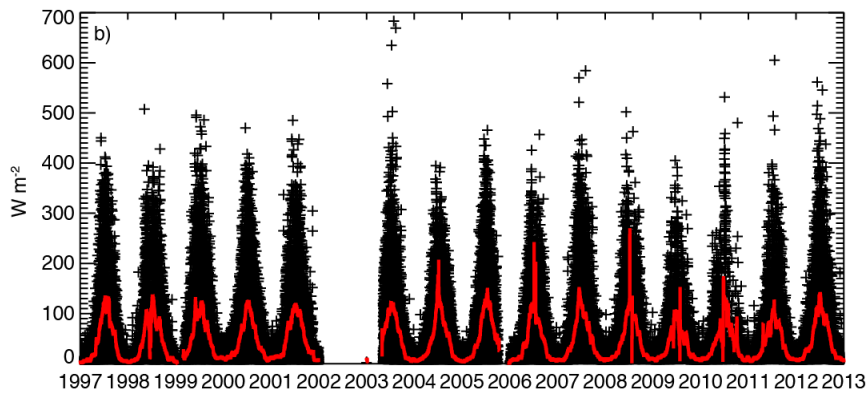
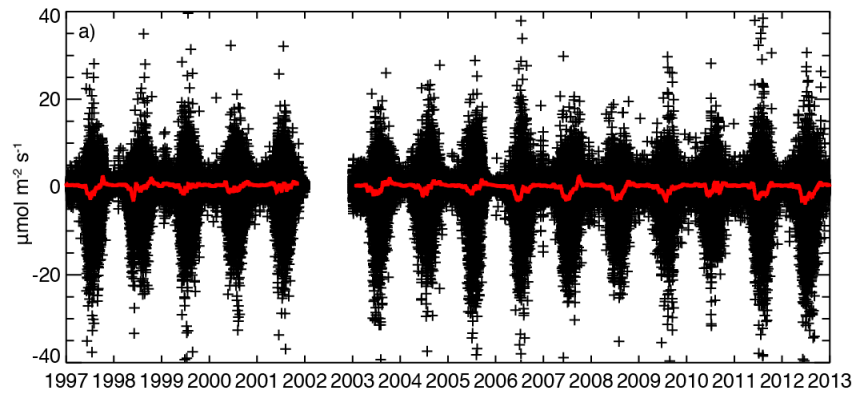
805 **Figures**

Willow Creek - NetCam SC IR - Thu Sep 20 11:31:17 2012
Temperature: 36.0 °C internal, 9.0 °C outside
RH: 0%, Pressure: 944.0 millibars
Exposure: 400

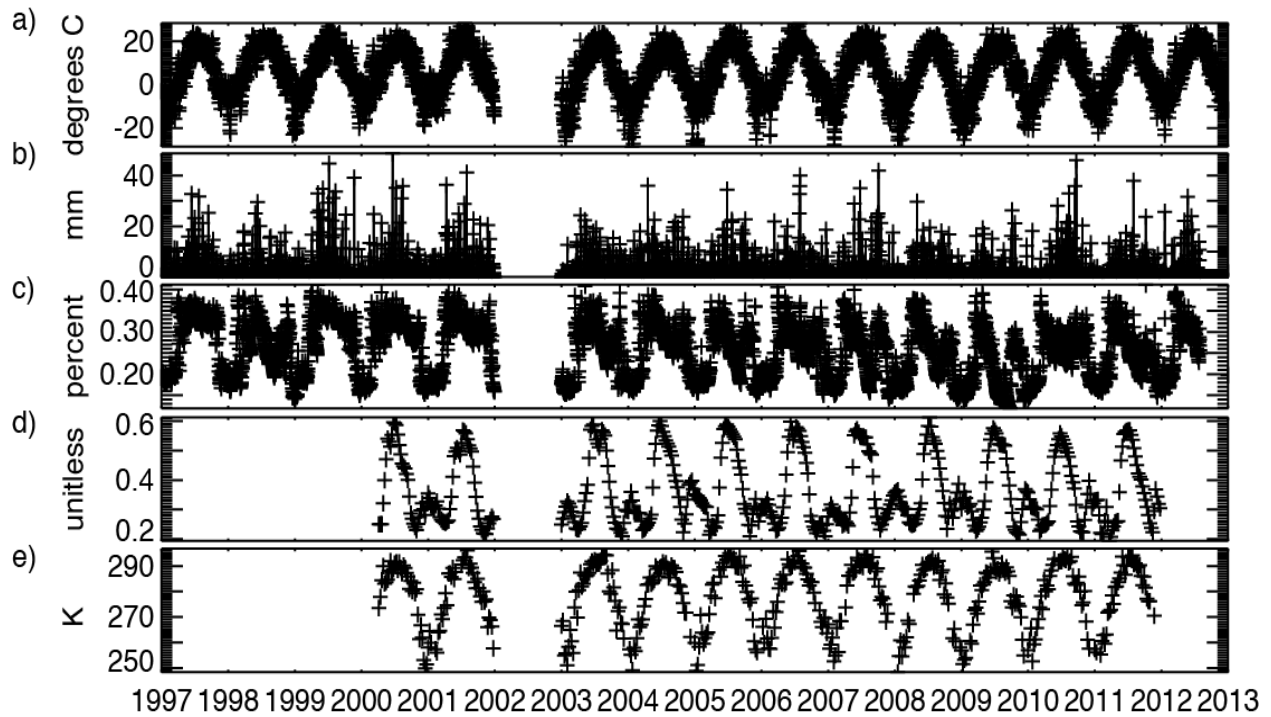


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Fig. 1 Examples of processes that affect photosynthesis in the environment that make canopy-scale photosynthesis differ from leaf-level photosynthesis, superimposed on a photo taken from an eddy covariance flux tower overlooking a temperate hardwood forest .

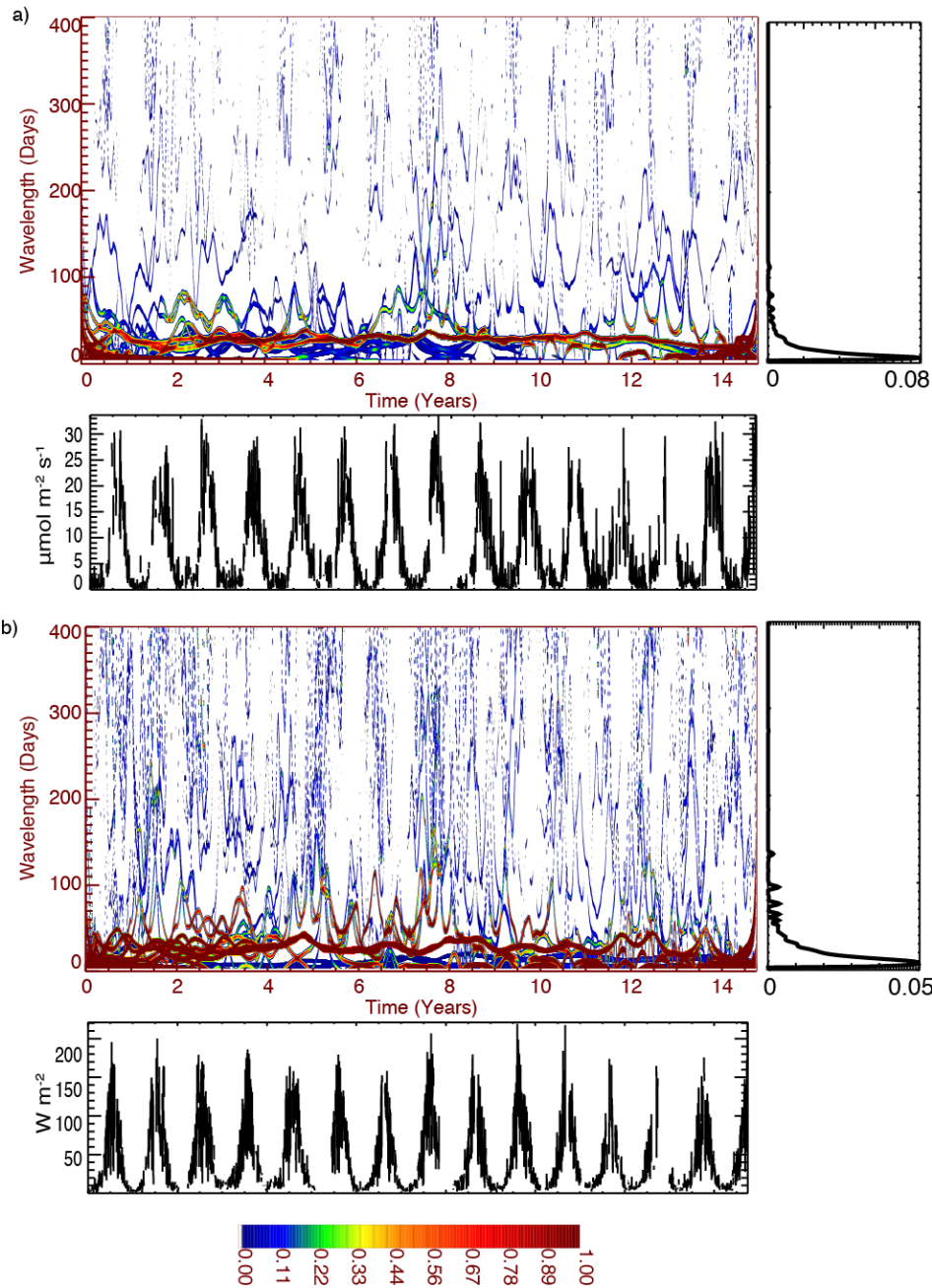


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 812 **Fig. 2** Nearly fifteen year time series of (a) net ecosystem exchange of CO₂ (*NEE*) and (b)
 813 evapotranspiration flux (*ET*) at hourly (black crosses) and biweekly (red line) timescales as
 814 observed from the WLEF Park Falls, WI very tall eddy covariance flux tower from 1997-
 815 2012. Long-term datasets like these on canopy-scale carbon and water fluxes are starting
 816 to be made available in the eddy covariance community. The large variability at multiple
 817 timescales presents a significant opportunity and challenge for improving predictions of
 818 ecosystem fluxes.
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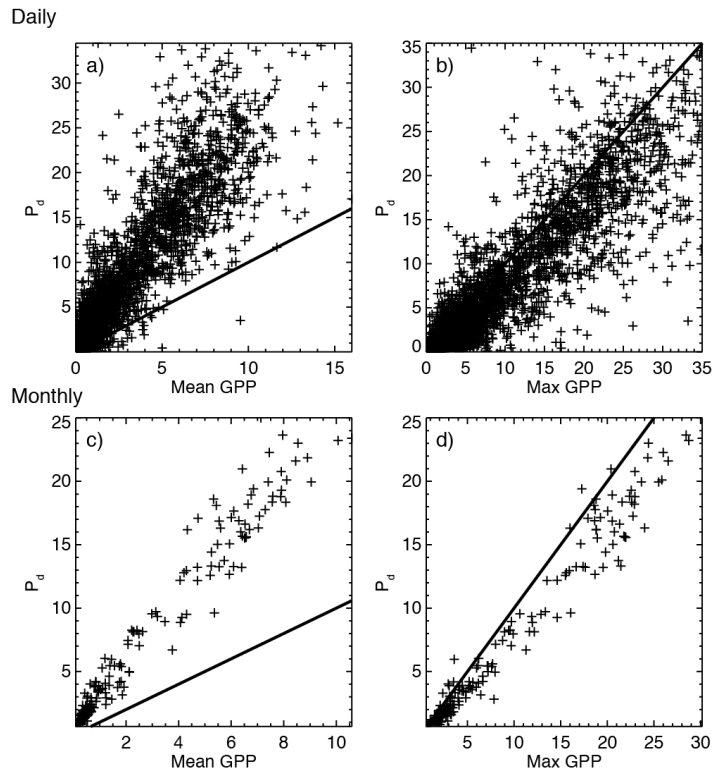
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Fig 3. Time series of daily a) air temperature (T_{air}), b) precipitation (P_{precip}), c) soil moisture (Q_{soil}), and biweekly d) MODIS enhanced vegetation index (EVI), and e) MODIS land surface temperature (LST). To allow comparisons to the quality-controlled flux data, 2002 was also removed from these data.



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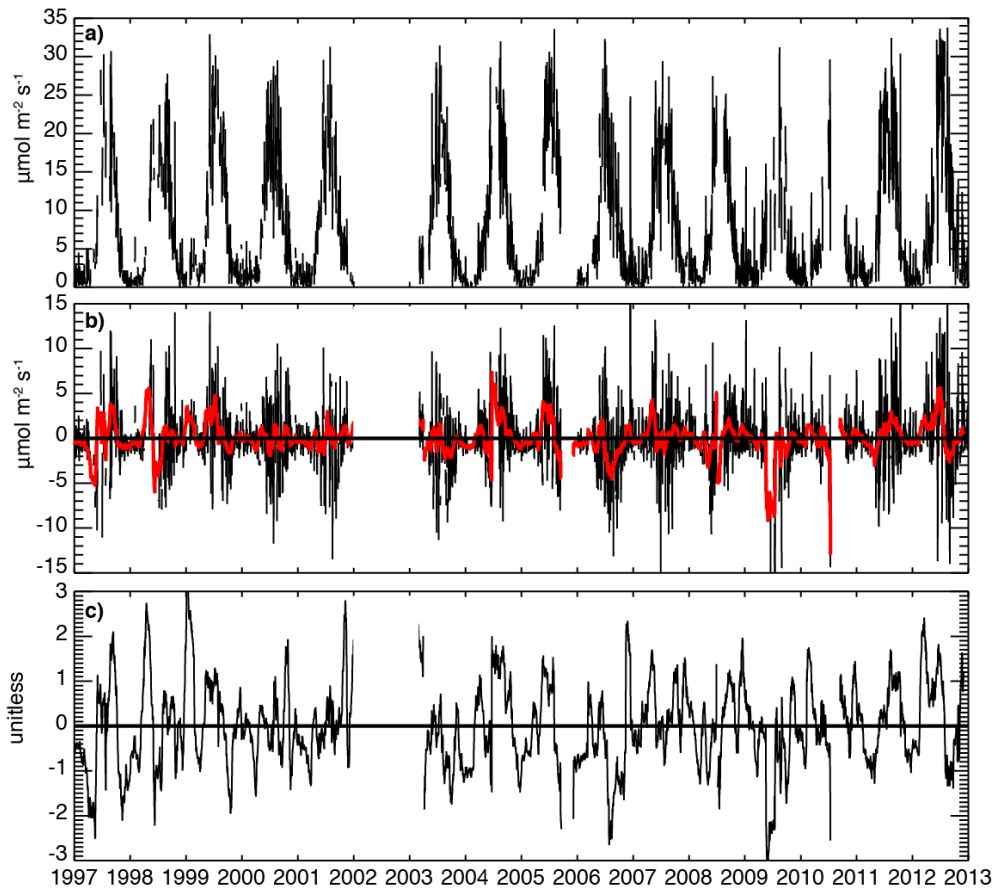
Fig. 4 Hilbert-Huang power spectral transformation of (a) net canopy carbon uptake (P_d) and (b) daily evapotranspiration (ET) derived from empirical mode decomposition of the tall tower fluxes. Raw daily time series are shown below each transformation and marginal power spectrum on the right. To improve clarity of higher frequency variation, the transform is cut-off at 400-day wavelength. The time-frequency decomposition reveals that both carbon and water exchange exhibit a number of characteristic timescales of variability, including substantial peaks at the monthly and seasonal timescale. Longer-term variability is much weaker, especially for carbon, but growing season signals of variability on the 100-200 day wavelength exist and strong coherence among water and carbon are also noted.



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Fig. 5 Scatterplot of flux tower derived GPP (based on fit of nighttime NEE to temperature to estimate respiration) to P_d for a) mean daily GPP, b) maximum daily GPP, c) mean monthly GPP, and d) mean monthly maximum daily GPP. 1:1 fit line shown on all four and shows that P_d is closely related to maximum daily GPP. Correlation at daily scale is $r^2=0.81$ and monthly $r^2=0.96$.

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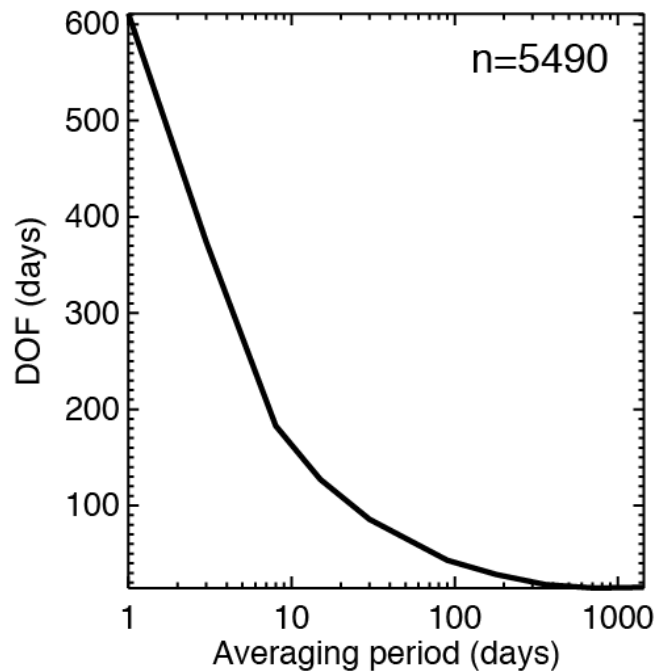
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850 **Fig. 6** Example of generating relative anomalies for lag correlation analysis. Raw daily time
851 series of (a) net photosynthetic drawdown (P_d), were (b) de-seasonalized by removal of the
852 ensemble average daily time series and then averaged to the appropriate averaging
853 timescale, in this example, monthly (red line) and finally (c) normalized to relative values.
854 The final signal represents the true anomalies of variation across time and has successfully
855 removed the seasonal variability of solar forcing.

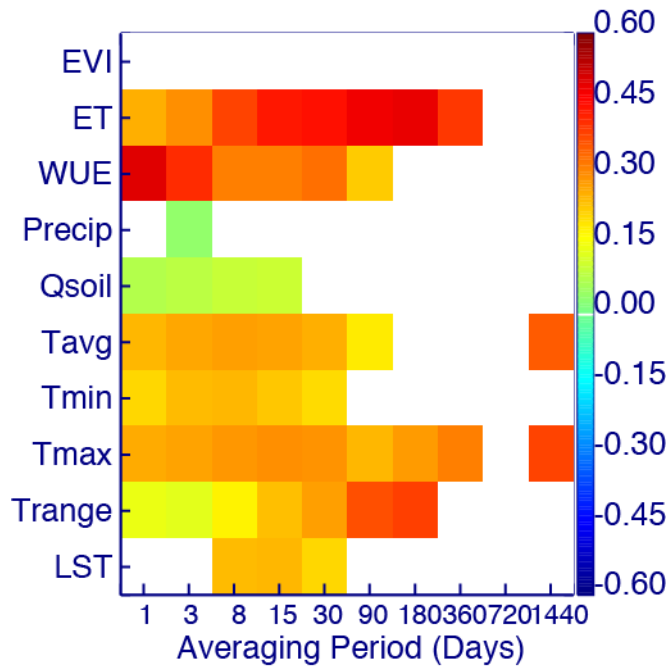
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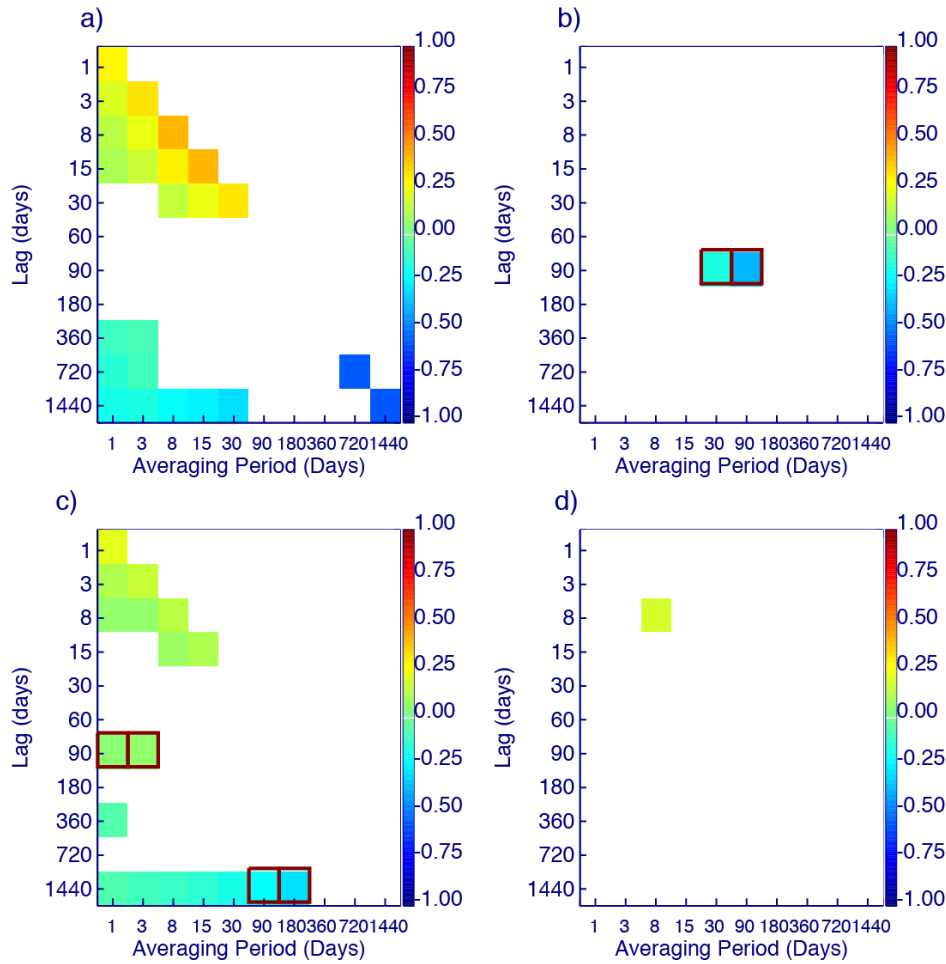
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Fig. 7 Degree of freedom analysis for daily net photosynthetic drawdown from the tall tower fluxes. The total record has 5490 observation periods (with 37% of observations missing), but the full autocorrelation analysis reveals an exponentially declining true degree of freedom from slightly over 600 for no-averaging, to near zero above 360 day smoothing of the time series. This reduced N is used for all correlation significance tests.



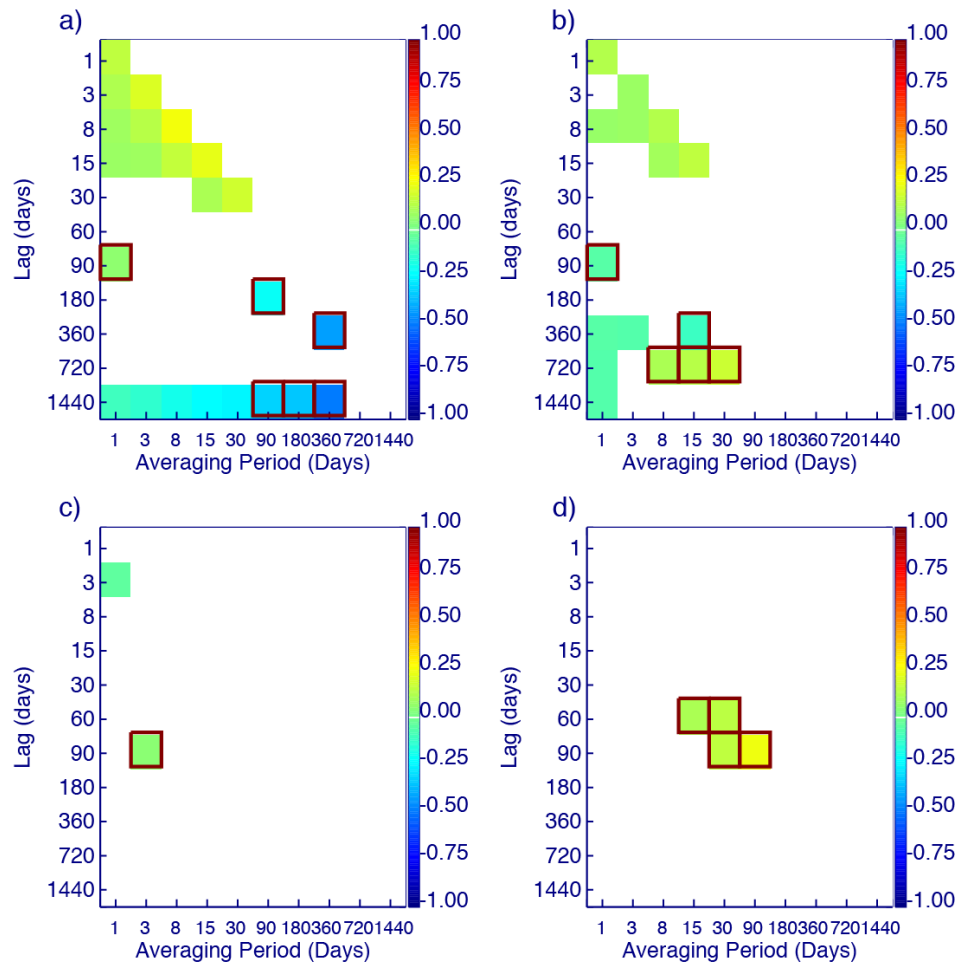
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Fig. 8 Direct correlation coefficients of standardized anomalies of climate variables to net photosynthetic drawdown (P_d) anomalies at daily to multi-year filter scales. Temperature variables explain a small fraction of weekly to monthly anomalies in P_d , while moisture variability explains a greater fraction of long-term seasonal variability in anomalies of P_d . Interestingly, remotely sensed vegetation index anomalies do not significantly explain P_d anomalies at any timescale.



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878 **Fig. 9** Analysis of lagged standardized anomalies of (a) P_d , (b) EVI , (c) T_{mean} , and (d) LST to
 879 P_d at a range of averaging times and lag times. Only significant correlations are displayed,
 880 after correction of autocorrelated degrees of freedom. Red squares indicate significant
 881 correlations that exceed P_d autocorrelation at that averaging and lag scale. Strong AR-1
 882 correlation persist in P_d at lags up to one month and averaging periods to one month, while
 883 small negative correlations show predictive ability for temperature at lags exceeding one
 884 year and EVI at the two-month lag.
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Fig. 10 Same as Fig. 9 but for moisture variables of (a) ET, (b) WUE, (c) P_{precip} , and (d) Q_{soil} . While ET and WUE have significant direct correlations as shown in Fig. 5, lagged correlations are small. Stronger positive correlations exist for precipitation and soil moisture at the weekly to seasonal timescale for seasonal scale lags.

892 **Tables**

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894 **Table 1** Abbreviations used in this paper

Abbreviation	Description	Source
<i>NEE</i>	Net ecosystem exchange of CO ₂	Flux tower
<i>GPP</i>	Gross primary production	Flux tower
<i>P_d</i>	Photosynthetic drawdown	Flux tower
<i>EVI</i>	Enhanced Vegetation Index, 8-day average	MODIS TERRA/AQUA
<i>ET</i>	Evapotranspiration	Flux tower
<i>WUE</i>	Water Use Efficiency (P_d/ET)	Flux tower
<i>P_{precip}</i>	Daily precipitation	NCDC + NARR Reanalysis
<i>Q_{soil}</i>	10 cm soil moisture	NARR Reanalysis
<i>T_{mean}</i>	Daily temperature	Flux tower + NCDC
<i>T_{min}</i>	Minimum daily temperature	Flux tower + NCDC
<i>T_{max}</i>	Maximum daily temperature	Flux tower + NCDC
<i>T_{range}</i>	Daily temperature range (max - min)	Flux tower + NCDC
<i>LST</i>	Land Surface Temperature, 8-day day/night average	MODIS TERRA/AQUA

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897 **Table 2** Granger causality analysis for P_d as a function of averaging period reveals best
 898 predictive models generally include temperature and transpiration observations for daily
 899 to weekly averages but precipitation, soil moisture, and water use efficiency for monthly
 900 time scales, primarily within the first or two lags. No variables could exceed P_d
 901 autoregression in explaining longer averaging scales (> 30 days).

Variable/Averaging period (Days)	1	3	8	15	30	90	360
<i>EVI</i>						90	
<i>T_{mean}</i>	1	3	8				
<i>LST</i>	1	3					
<i>ET</i>	1-3	3	8				360
<i>WUE</i>	1-3	3		15-360			
<i>P_{precip}</i>							
<i>Q_{soil}</i>							

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